# COURSE TITLE: ECONOMETRICS 

DEPARTMENT: AGRICULTURAL ECONOMICS \&<br>FARM MANAGEMENT

## LECTURER: DR S.A ADEWUYI

- Meaning and concepts of Econometrics
- Correlation analysis
- Simple regression analysis
- Multiple regression analysis
- Problems associated with the regression analysis
- Functional forms.
- MEANING AND CONCEPT OF ECONOMETRICS

Econometrics is the scientific application of statistics and mathematics to economic theory.

Steps in Econometrics

1. Selection of variables
2. Construction of mathematical model
3. Data collection- cross-sectional data, time series data and pooled data
4. Estimation of data
5. Data estimation.

Criteria for evaluating econometric data

1. Economic criterion
2. Statistics criterion
3. Econometric criterion

## SIMPLE CORRELATION

Correlation is a measure of association that exists between 2 or more variables. Correlation analysis attempts to find out the degree or extent to which variables tend
to move together. For instance, in demand theory. Correlation analysis is used to show the degree of relationship between the prize and the quantity.

Types of correlation

1) Positive, negative and zero
2) Simple, partial and multiple
3) Linear and Non-Linear

If ' $r$ ' is the correlation coefficient between 2 variables $X$ and $Y$ then.

'r' always lies between -1 and +1 . Correlation may be positive, negative or zero. When $\mathrm{r}=1$, the 2 variables are said to be positively correlated. When $\mathrm{r}=-1$, they are said to be negatively correlated. When $\mathrm{r}=\mathrm{o}$. there is no correlation between the 2 variables.

## Positive Correlation.

If two variables are positively Correlated, their values tend to rise fall together.


As shown in the graph, when X increases, Y also increases implies that is an agreement between X and Y .

## Negative Correlation

Under negative Correlation, 2 variables more in opposite direction. As one increases, the other one decreases i.e there is no agreement between them. This is depicted in the graph below.


## Zero Correlation

This occurs when there is no joint movement between two variables. A situation of zero correlation between X and Y is shown below.


## Zero Correlation

The diagram is referred to as the scattered diagram.
The correlation coefficient can also be obtained as follows:

1. Pearson (product - moment) correlation coefficient.
(a) Using the standard deviation method.

$$
\begin{aligned}
& \underset{\text { r Sxy }}{\text { Sx Sy }} \quad ; \operatorname{Sxy}=\sum(\mathrm{X}-\mathrm{X})(\overline{\mathrm{Y}}-\mathrm{Y}) \overline{=} \text { Standards deviation of } \mathrm{X} \text { and } \mathrm{Y} \\
& \mathrm{Sx}=\sqrt{\frac{\mathrm{N}}{\frac{\sum(\mathrm{X}-\mathrm{X})^{2}}{\mathrm{~N}}} \quad=\text { Standard deviation of } \mathrm{X} .} \\
& \mathrm{Sy}=\sqrt{\frac{\sum(\mathrm{Y}-\mathrm{Y})^{2}}{} \quad=\text { Standards deviation of } \mathrm{Y} .}
\end{aligned}
$$

## N

. . The correlation coefficient ' $r$ ' is given as


## SIMPLE LINEAR REGRESSION

Regression is the amount of change in value of one variable associated with a unit change in the value o another variable. It shows the dependence of a random variable, on another variable $X$, which is not necessarily a random variable; an equation, which relates Y to X , is usually called a regression equation. Two variables are involved in
regression analysis. They are the dependent or the endogenous variable and the independent or the exogenous variable.

Regression analysis is simply when only two variable (one independent and dependent) are involved. It is linear in the sense that we want to fit a straight line to a set of data involving the 2 variables.

## Assumptions of the simple linear model

1. The regression model of the linear in the unknown coefficient
2. The explanatory variables are not perfectly linearly correlated.
3. The error term $\left(\mathrm{e}_{\mathrm{i}}\right)$ is a random real variable
4. The mean of $e_{i}$ in any particular period is zero i.e. $E\left(e_{i}\right)$
5. The variable of e 1 is constant in each period, i.e. $\operatorname{Var}\left(\mathrm{e}_{\mathrm{i}}{ }^{2}\right)=S^{2}$
6. The variance has a normal distribution
7. The explanatory variance are measured without any error
8. The error term $\left(e_{i}\right)$ is normally distributed so that $U_{i}---N\left(O, \&{ }^{2}\right)$
9. Not all independent variables $\left(\mathrm{X}_{\mathrm{s}}\right)$ are the same; at least one of them is different.
10.The random term of different observation $\left(\mathrm{U}_{\mathrm{i}}, \mathrm{U}_{\mathrm{j}}\right)$ are independently distributed so that $\quad \operatorname{Cov}\left(\mathrm{U}_{\mathrm{i}}, \mathrm{U}_{\mathrm{j}}\right)=\mathrm{E}\left(\mathrm{U}_{\mathrm{i}}, \mathrm{U}_{\mathrm{j}}\right)=0$

Suffice is to say, all these assumptions must hold before we use the ordinary least square model.

## Normal equation and the estimation of regression parameters

Given a regression line $Y_{i}=a+b_{i} X_{1}+e_{i}$

Equations 3 and 4 are normally least square equations. From equation 3

$$
\mathrm{b}_{0}=\quad \underset{/ \mathrm{n}}{\sum \mathrm{v}_{i,}} \quad \frac{\mathrm{~b}_{\mathrm{i}} \sum \mathrm{x}_{\mathrm{i}}}{\mathrm{n}}
$$

$$
\mathrm{b}_{0}=\mathrm{y}-\mathrm{b}_{\mathrm{i}} \mathrm{x}_{\mathrm{i}-\ldots}
$$

Substituting equation (5) in (4)

$$
\begin{aligned}
& \sum \mathrm{x}_{\mathrm{i}} \mathrm{y}=\left(\mathrm{y}-\mathrm{b}_{\mathrm{i}} \mathrm{x}\right) \sum \mathrm{x}_{\mathrm{i}}+\mathrm{b}_{\mathrm{i}} \mathrm{x}_{\mathrm{i}}^{2} \\
& \sum \mathrm{x}_{\mathrm{i}} \mathrm{y}=\left(\sum \mathrm{x}_{\mathrm{i}}-\mathrm{b}_{\mathrm{i}} \mathrm{x} \sum \mathrm{x}_{\mathrm{i}}+\mathrm{b}_{\mathrm{i}} \sum \mathrm{x}_{\mathrm{i}}^{2}\right.
\end{aligned}
$$

$$
\begin{align*}
& \mathrm{ei}=\mathrm{Yi}-\mathrm{b}_{\mathrm{o}}-\mathrm{b}_{\mathrm{o}} \mathrm{Xi} \\
& \text { Min } \quad \sum \mathrm{e}_{\mathrm{i}}^{2}=\min \sum\left(\mathrm{Y}_{\mathrm{i}}-\mathrm{b}_{\mathrm{o}}-\mathrm{b}_{0} \mathrm{Xi}\right)^{2} \\
& \frac{\partial \sum \mathrm{e}_{\mathrm{i}}{ }^{2}}{\partial \mathrm{bo}}=-2\left(\mathrm{Y}_{\mathrm{i}}-\mathrm{b}_{\mathrm{o}}-\mathrm{b}_{\mathrm{i}} \mathrm{Xi}\right)=0 \\
& \partial \sum \mathrm{e}_{\mathrm{i}}^{2}=-2\left(\mathrm{Y}_{\mathrm{i}}-\mathrm{b}_{\mathrm{o}}-\mathrm{b}_{\mathrm{i}} \mathrm{Xi}\right)=0 \\
& \partial b_{i} \\
& \sum \mathrm{y}_{\mathrm{i}}-\sum \mathrm{b}_{0}-\mathrm{b}_{\mathrm{i}} \sum \mathbf{X}_{\mathrm{ii}}=0  \tag{1}\\
& \sum_{\mathrm{xi}} \mathrm{y}_{\mathrm{i}}=\mathrm{b}_{0} \sum \mathbf{X}_{\mathrm{i}}-\mathrm{b}_{1} \sum \mathrm{x}_{\mathrm{i}}{ }^{2}=  \tag{2}\\
& \sum \mathrm{y}_{\mathrm{i}}=\sum \mathrm{b}_{0}+\mathrm{b}_{1} \sum \mathrm{x}_{\mathrm{i}} \\
& \sum_{\mathrm{xi}} \mathrm{y}_{\mathrm{i}}=\mathrm{b}_{0} \sum \mathrm{x}_{\mathrm{i}}+\mathrm{b}_{1} \sum \mathrm{x}_{\mathrm{i}}{ }^{2} \\
& \sum \mathrm{y}_{\mathrm{i}}=\mathrm{nb}_{0}+\mathrm{b}_{\mathrm{i}} \sum \mathrm{x}_{\mathrm{i}} \tag{3}
\end{align*}
$$

$$
\begin{aligned}
& \mathrm{b}_{\mathrm{i}}= \frac{\sum \mathrm{x}_{\mathrm{i}} \mathrm{y}_{\mathrm{i}}-\mathrm{y} \sum \mathrm{x}_{\mathrm{i}}}{\sum \mathrm{x}_{\mathrm{i}}^{2} \mathrm{y}_{\mathrm{i}}-\mathrm{y} \sum \mathrm{x}_{\mathrm{i}}} \\
&= \sum \mathrm{x}_{\mathrm{i}} \mathrm{y}_{\mathrm{i}}-\sum \mathrm{x}_{\mathrm{i}} \sum \mathrm{y}_{\mathrm{i}} / \mathrm{n} \\
& \sum \mathrm{x}_{\mathrm{i}}= \sum \mathrm{x}_{\mathrm{i}} \sum \mathrm{x}_{\mathrm{i}} / \mathrm{n} \\
& \mathrm{n} \sum \mathrm{x}_{\mathrm{i}} \mathrm{y}_{\mathrm{i}}-\sum \mathrm{x}_{\mathrm{i}} \mathrm{y}_{\mathrm{i}} \\
& \mathrm{~b}_{\mathrm{i}}= \frac{\mathrm{n} \sum \mathrm{x}_{\mathrm{i}}^{2}-\sum \mathrm{x}_{\mathrm{i}}^{2}}{\sum \mathrm{x}_{\mathrm{i}} \mathrm{y}_{\mathrm{i}}} \\
& \sum \mathrm{x}_{\mathrm{i}}^{2}
\end{aligned}
$$

## Example

The Table below shows the respective salaries of a father
(X) and his oldest Son (Y) of a sample of 12

(a) Find the least square regression line of $Y$ on $X$
(b) Estimate the salary of the oldest son if the salary of the father is N9,500

Solution:

| $\mathbf{X}$ | $\mathbf{Y}$ | $\mathbf{X}^{\mathbf{2}}$ | $\mathbf{X Y}$ |
| :--- | :--- | :--- | :--- |
| 5 | 8 | 25 | 40 |
| 3 | 6 | 9 | 18 |
| 7 | 8 | 49 | 56 |


| 4 | 5 | 16 | 20 |
| :--- | :--- | :--- | :--- |
| 8 | 9 | 64 | 72 |
| 2 | 6 | 4 | 12 |
| 10 | 8 | 100 | 80 |
| 6 | 5 | 36 | 30 |
| 8 | 11 | 64 | 88 |
| 7 | 7 | 49 | 49 |
| 9 | 8 | 81 | 72 |
| 11 | 10 | 121 | 110 |

$\sum \mathrm{X}=80, \quad \sum \mathrm{Y}=91, \quad \sum \mathrm{XY}=647, \quad \sum \mathrm{X} 2=618$
Let the regression line of Y on X be

$$
Y=a+b X
$$

$$
B=\frac{N \sum X y-\sum X \sum Y}{N \sum X 2-\left(\sum X\right)^{2}}
$$

$$
=\quad 12 \times 647-80 \times 91
$$

$$
12 \times 618-(80)^{2}
$$

$$
=\quad 7764-7280
$$

$$
7416-6400
$$

$$
=\quad 484
$$

$$
1016
$$

$$
=0.4764
$$

a

$$
\begin{aligned}
& =Y-b_{i} X \\
& =\frac{\sum Y}{N}-\frac{b \sum X}{12}=\frac{0.4764}{12} \mathrm{X} 80 \\
& =\frac{91}{12} \\
& =9.5833-3.176 \\
& =4.4073
\end{aligned}
$$

Substituting a and b in the equation
$Y=a+b_{i} X$
$\mathrm{Y}=4.4073+0.4764 \mathrm{X}$
This is the required regression line.
(b). If $X=\# 9,500$, from the equation

$$
\begin{aligned}
Y & =4.4073+0.4764 \mathrm{X} \\
& =4.4073+0.4764(9520) \\
& =453.20
\end{aligned}
$$

The regression line shows that there is a positive relationship between the salary of the father $(\mathrm{X})$ and that of the oldest son $(\mathrm{Y})$. It further implies that if the salary of the son increases by \#1, the father's salary will also increase by \#0.48.

## MULTIPLE REGRESSION

Regression can be defined as the amount of change in the value of one variable associated with a unit change in the value of another variable. There are two types of variables involved in regression analysis. These are independent and dependent variables. Multiple regressions analysis used involves three or more
variables.
Multiple regression analysis is used for testing the relationship between dependent variable, Y and two or more independent variables, Xs , and for prediction.

The variable linear regression models can be written as
$\mathrm{Y}=\mathrm{b}_{0}+\mathrm{b}_{\mathrm{i}} \mathrm{X}_{\mathrm{i}}+\mathrm{b}_{2} \mathrm{X}_{2}$
equation 6
Ordinary least square parameter estimates for equation can be obtained by minimizing the sum of the square residuals
$\sum \mathrm{ei}^{2}=\sum(\mathrm{Y}-\hat{\mathrm{Y}})^{2}=\left(\mathrm{Y}-\mathrm{b}_{\mathrm{o}}+\mathrm{b}_{1} \mathrm{X}_{1}+\mathrm{b}_{2} \mathrm{X}_{2}\right)^{2}$ equation 7

The normal equations are.
$\sum \mathrm{Y}=\mathrm{b}_{\mathrm{o}}+\mathrm{b}_{1} \sum \mathrm{X}_{1}+\mathrm{b}_{2} \sum \mathrm{X}_{2}$
$\sum \mathrm{X}_{1} \mathrm{Y}=\mathrm{b}_{\mathrm{o}} \sum \mathrm{X}_{1}^{2}+\mathrm{b}_{2} \sum \mathrm{X}_{1} \mathrm{X}_{2}$ equation 8
$\sum \mathrm{X}_{2} \mathrm{Y}=\mathrm{b}_{\mathrm{o}} \sum \mathrm{X}_{1}^{2}+\mathrm{b}_{1} \sum \mathrm{X}_{1} \mathrm{X}_{2}+\mathrm{b}_{2} \sum \mathrm{X}_{2}$
In deviation for the normal equations become
$\sum \mathrm{X}_{1} \mathrm{Y}=\mathrm{b}_{\mathrm{o}} \sum \mathrm{X}_{1}^{2}+\mathrm{b}_{2} \sum \mathrm{X}_{1} \mathrm{X}_{2}$
$\sum \mathrm{X}_{2} \mathrm{Y}=\mathrm{b}_{\mathrm{o}} \sum \mathrm{X}_{2}^{2}+\mathrm{b}_{1} \sum \mathrm{X}_{1} \mathrm{X}_{2}+\mathrm{b}_{2} \sum \mathrm{X}_{2}^{2}$ equation 9
$\sum \mathrm{X}_{1}{ }^{2}=\sum \mathrm{X}_{1}{ }^{2}-\overline{\mathrm{n}}(\mathrm{X})^{2}$
$\sum \mathrm{X}_{1} \mathrm{Y}=\sum \mathrm{X}_{1} \mathrm{Y}-\overline{\mathrm{n}}(\mathrm{Y})^{2}$
$\sum \mathrm{X}_{1} \mathrm{Y}^{2}=\sum \mathrm{X}_{1} \mathrm{Y}^{2}-\mathrm{n} \mathrm{X}_{1} \mathrm{Y}$
$\sum \mathrm{X}_{2}{ }^{2}=\sum \mathrm{X}_{2}{ }^{2}-\mathrm{n}\left(\mathrm{X}_{2}\right)^{2}$
$\sum X_{2} Y=\sum X_{2} Y-\bar{n} \bar{X}_{2} Y$
$\left(\begin{array}{ll}\sum \mathrm{X}_{1}{ }^{2} & \sum \mathrm{X}_{1} \mathrm{X}_{2} \\ \sum \mathrm{X}_{1} \mathrm{X}_{2} & \sum \mathrm{X}_{2}{ }^{2}\end{array}\right)\binom{\mathrm{b}_{1}}{\mathrm{~b}_{2}}=\binom{\sum \mathrm{X}_{1} \mathrm{Y}}{\sum \mathrm{X}_{2} \mathrm{Y}}$

$$
\binom{b_{1}}{b_{2}}=\left(\begin{array}{ll}
\sum X_{1}^{2} & \sum X_{1} X_{2} \\
\sum X_{1} X_{2} & \sum X_{1}^{2}
\end{array}\right)^{-1} \quad\binom{\sum X_{1} Y}{\sum X_{2} Y}
$$

## Original Matrix

$\left(\begin{array}{ll}\sum \mathrm{X}_{1}{ }^{2} & \sum \mathrm{X}_{1} \mathrm{X}_{2} \\ \sum \mathrm{X}_{1} \mathrm{X}_{2} & \sum \mathrm{X}_{2}{ }^{2}\end{array}\right)$

## Matrix of Minor

$$
\left(\begin{array}{ll}
\sum \mathrm{X}_{2}^{2} & \sum \mathrm{X}_{1} \mathrm{X}_{2} \\
\sum \mathrm{X}_{1} \mathrm{X}_{2} & \sum \mathrm{X}_{1}^{2}
\end{array}\right)
$$

## Cofactor

$$
\left(\begin{array}{ll}
\sum \mathrm{X}_{2}^{2} & -\sum \mathrm{X}_{1} \mathrm{X}_{2} \\
\sum \mathrm{X}_{1} \mathrm{X}_{2} & +\sum \mathrm{X}_{1}^{2}
\end{array}\right)
$$

## Transpose

$\sum \mathrm{X}_{2}{ }^{2}-\sum \mathrm{X}_{1} \mathrm{X}_{2}$
$\sum \mathrm{X}_{1} \mathrm{X}_{2} \quad \sum \mathrm{X}_{1}{ }^{2}$

## Determinant

$$
\begin{aligned}
& \sum \mathrm{X}_{1}{ }^{2} \mathrm{X}_{2}^{2}-\sum \mathrm{X}_{1} \mathrm{X}_{2}{ }^{2} \\
& \text { Inverse }=\text { Transpose }
\end{aligned}
$$

$$
\begin{gathered}
=\left(\begin{array}{c}
\text { Determinant } \\
\sum \mathrm{X}_{2}{ }^{2}-\sum \mathrm{X}_{1} \mathrm{X}_{2} \\
\frac{-\sum \mathrm{X}_{1} \mathrm{X}_{2} \sum \mathrm{X}_{1}{ }^{2}}{\sum \mathrm{X}_{1}{ }^{2} \mathrm{X}_{2}{ }^{2}-\sum \mathrm{X}_{1} \mathrm{X}_{2}{ }^{2}}
\end{array}\right)
\end{gathered}
$$

$$
\begin{aligned}
& b_{1}=\left(\begin{array}{cc}
\sum \mathrm{X}_{2}{ }^{2}-\sum \mathrm{X}_{1} \mathrm{X}_{2} \\
\mathrm{~b}_{2}=\left(\begin{array}{ll}
-\sum \mathrm{X}_{1} \mathrm{X}_{2} \sum \mathrm{X}_{1}{ }^{2} & \sum \mathrm{X}_{2} \mathrm{Y}
\end{array}\right. \\
\sum \mathrm{X}_{1}{ }^{2} \sum \mathrm{X}_{2}{ }^{2}-\left(\sum \mathrm{X}_{1} \mathrm{X}_{2}\right)^{2}
\end{array}\right) \sum \mathrm{X}_{1} \mathrm{Y} \\
& \mathrm{~b}_{1}=\sum \mathrm{X}_{2}^{2} \sum \mathrm{X}_{1}{ }^{2} \mathrm{Y}-\sum \mathrm{X}_{1}{ }^{2} \mathrm{X}_{2} \sum \mathrm{X}_{1}{ }^{2} \mathrm{Y} \\
& \sum \mathrm{X}_{1}^{2} \sum \mathrm{X}_{2}^{2}-\left(\mathrm{X}_{1} \sum \mathrm{X}_{2}^{2}\right)^{2} \\
& \mathrm{~b}_{2}=\frac{\sum \mathrm{X}_{2} \mathrm{Y} \sum \mathrm{X}_{1}{ }^{2}-\sum \mathrm{X}_{1} \mathrm{Y} \sum \mathrm{X}_{1} \mathrm{X}_{2}}{\sum \mathrm{X}_{1}{ }^{2} \sum \mathrm{X}_{2}{ }^{2}-\left(\mathrm{X}_{1} \sum \mathrm{X}_{2}{ }^{2}\right)^{2}} \\
& \text { Во }=\overline{\mathrm{Y}}-\overline{\mathrm{b}}_{1} \mathrm{X}_{1}-\overline{\mathrm{b}}_{2} \mathrm{X}_{2} \\
& \underline{\text { Variance of error term ( } \mathbf{S}^{2} \text { ) }}
\end{aligned}
$$

$$
\mathrm{S}^{2}=\frac{\sum \mathrm{e}^{2}}{\mathrm{n}-\mathrm{k}} \frac{=\sum \mathrm{Y}^{2}-\left(\mathrm{bo}\left(\sum \mathrm{Y}+\mathrm{b}_{1} \sum \mathrm{X}_{1} \mathrm{Y}+\mathrm{b}_{2} \sum \mathrm{X}_{2} \mathrm{Y}\right)\right.}{\mathrm{n}-\mathrm{k}}
$$

Where $S^{2}=$ Variance of error
n = Sample Size
$K=$ No of parameters

## Variance of coefficients

$\operatorname{Var}\binom{b_{1}}{b_{2}}=S^{2}\left(X^{1} X\right)^{-1}$
$\operatorname{Var}\binom{\mathrm{b}_{1}}{\mathrm{~b}_{2}}=\left(\begin{array}{ccc}\mathrm{S}^{2} & \sum \mathrm{X}_{2}^{2} \mathrm{X} & \sum \mathrm{X}_{1} \mathrm{X}_{2} \\ & \sum \mathrm{X}_{1} \mathrm{X}_{2} & \sum \mathrm{X}_{1}{ }^{2}\end{array}\right)^{-1}$

The variances of b1 and b2 are found in the principal diagonal of the matrix starting with $\operatorname{var}(\mathrm{bi})$ in the upper left corner and ending with $\operatorname{var}(\mathrm{b} 2)$ in the lower right corner Therefore
$\operatorname{Var}(\mathrm{bi})=S^{2}\left(\frac{\sum X_{2}{ }^{2}}{\sum \mathrm{X}_{1}{ }^{2} \mathrm{X}_{2}{ }^{2}-\sum \mathrm{X}_{1} \mathrm{X}_{2}{ }^{2}}\right)$ equation 15
$\operatorname{Var}(\mathrm{b} 2)=\mathrm{S}^{2}\left(\frac{\sum \mathrm{X}_{1}{ }^{2}}{\sum \mathrm{X}_{1}{ }^{2} \mathrm{X}_{2}{ }^{2}-\left(\sum \mathrm{X}_{1} \mathrm{X}_{2}\right)^{2}}\right)$
equation 16

## Coefficient of determination $\left(\mathbf{R}^{\mathbf{2}}\right)$

This is used t measure how well the estimated regression equation fits the observed data on X and Y . The coefficient measures the "goodness of fit"


## Adjusted Coefficient of determination (R2)

$\mathrm{R}^{-2}=\left(1-\left(1-\mathrm{R}^{2}\right) \frac{\mathrm{n}-1}{\mathrm{n}-\mathrm{k}}\right.$

## PROBLEMS OF SINGLE EQUATION REGRESSION MODEL

1. Autocorrelation: The problems are arises when the assumption that error are independently fails to hold.

$$
\operatorname{Cov}\left(\mathrm{e}_{\mathrm{i}}, \mathrm{e}_{\mathrm{j}}\right)=0
$$

The first order autocorrelation can be obtained from simple residual error as:

$$
\text { et }=\beta e_{t}-1+v_{t}
$$

Where:
$\beta=$ Auto residual error
$e_{t}=$ Autocorrelation coefficient
$\mathrm{v}_{\mathrm{t}}=$ Random autocorrelation error.

TYPES OF AUTOCORRELATION

1. 1st Order
2. 2nd Order
3. Higher Order.

## CAUSES OF AUTOCORRELATION

1. Omission of important variable
2. Misspecification of function
3. Nature of economic data

## EFFECT OF AUTOCORRELATION

1. Large standard error
2. Universal estimators
3. Ordinary Least Square becomes inefficient.
4. Use of scattered diagram
5. Trial estimation of autocorrelation equation by testing the significance of the autocorrelation coefficient.
6. Use of Durbin Watson Statistics.

## MULTICOLLINEARITY

One of the basic assumptions of ordinary Least Square is that explanatory variables are independent. The violation of this assumption causes multicollinearity.
$E\left(X_{i} X_{j}\right)=0$

## SOURCES

1. Small Sample Size
2. Nature of economic data
3. Incorporation of lagged independent variable

## EFFECTS

1. Large standard error
2. Unbiased value of regression coefficient.

## METHODS OF DETECTION

1. Klein's rule
2. Beaton and Glamber method
3. Fairrar-Glauber Test
4. Uses of coefficient of determination $\left(\mathrm{R}^{2}\right)$
5. Use of Statistics

Heteroscedasticity occurs when the assumption of a mean value and a constant variance of the Ordinary Least Square fail to hold. When the assumption holds there is Heteroscedasticity. The problem of Heteroscedasticity is more common with cross sectional data than with time series data.

## CAUSES

1. Use of Heteroscedasticity Samples

## CONSEQUENCES

1. Regression parameters estimates are unbiased
2. Large variance of parameter estimate
3. Inefficient predicted estimates of the Least Square.

## CORRECTIONS

1. Weighed regression method
2. Use inverse Least Square Method
3. Method of instrumental variance

## STEPS IN ECOMOMETRIC STUDY

1. Selection of variable
2. Construction of mathematical model
3. Data collection viz: Cross sectional and pooled data.

## CRITERIA OF EVALUATING ECONOMETRIC DATA

1. Economic criterion
2. Statistical criterion
3. Econometric criterion.
4. Quadratic function

$$
\mathrm{Y}=\mathrm{b}_{\mathrm{o}}+\mathrm{b}_{\mathrm{i}} \mathrm{X}_{\mathrm{i}}+\mathrm{b}_{2} \mathrm{X}_{\mathrm{i}}^{2}+\mathrm{b}_{3} \mathrm{X}_{3}+\mathrm{b}_{4} \mathrm{X}_{2}^{2}
$$

Used for production and cost equations
2. Hyperbolic functions
$Y=b_{o}+b_{1} \frac{1}{X_{1}}+\frac{1}{b_{2} X_{2}}$
It is an isoquant equation.
3. Reciprocal function
$Y=\frac{1}{b_{0}+b_{i} X_{i}+b_{2} X_{2}}$
It is an isoquant equation
4. Square root function
$\mathrm{Y}=\mathrm{b}_{1} \mathrm{X}^{1 / 2}-\mathrm{b}_{2} \mathrm{X}_{2}^{1 / 2}$

Used for production and demand equation
5. Logarithm function

In $Y=\operatorname{Inb}_{\mathrm{o}} \mathrm{X}+$ bi $\operatorname{In} \mathrm{X}_{\mathrm{i}}+\mathrm{b}_{2} \operatorname{In} \mathrm{X}_{2}$
It is used for production function.

The regression coefficients reports represent elasticity coefficient s when the sum of the coefficient is $>1$, it implies decreasing return to scale.

When the sum of the coefficient $=1$, it implies constant return to scale.
6. Semi logarithmic function
$\mathrm{E}^{\mathrm{Y}}=\mathrm{boX}_{\mathrm{i}}{ }^{\mathrm{bi}} \mathrm{X}_{2}{ }^{\mathrm{b} 2}$
Used for cost and supply equation
7. Exponential function

$$
Y=e^{b o+b i X i+b 2 X 2+e}
$$

It is good for cost and supply equation.

## 8. Linear function

$$
Y=b_{o}+b_{i} X_{i}+b_{2} X_{2}+e
$$

Used for demand and supply function.

